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Shifting the Fat-Tailed Distribution of Blockbuster Punitive Damages Awards

by

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Abstract

The distribution of blockbuster punitive damages awards has fat tails similar to the distributions of losses from natural disasters. Extremely large awards occur more often and are more difficult to predict than if blockbuster awards were distributed normally. The size and predictability of awards are important factors in the U.S. Supreme Court's decisions on punitive damages. This article examines the effect of the Court's decision in *State Farm v. Campbell* on blockbuster punitive damages awards. *State Farm* shifts the fat tail of the distribution of blockbuster awards down (or "thins" the tail), which is consistent with a restraining effect on award size. *State Farm* reduces the size of blockbuster awards in general, but this reduction is most salient in the upper half of the distribution of awards. *State Farm* also has a negative influence on the probability of exceeding a single digit ratio between punitive and compensatory damages. This article also examines the largest awards and considers why defendants may not pay large punitive damages awards.

JEL codes: K10, K13, K40, K41

Keywords: punitive damages, blockbuster awards, fat tail, State Farm, product liability, tort

I. INTRODUCTION

Punitive damages share with natural disasters the property that there are often extremely large outliers that would not be consistent with a normal distribution. More specifically, extremely large punitive damages award are characterized by fat-tailed distributions. For natural disasters, these outliers take the form of extremely large losses. In the case of punitive damages, it is the magnitude of the awards that may be especially great.

In this article we focus specifically on what Viscusi (2004), Hersch and Viscusi (2004), and Del Rossi and Viscusi (2010) have termed blockbuster punitive damages awards, that is, punitive damages awards of at least \$100 million. Much like loses from significant natural disasters, these awards are extreme outliers and potentially pose catastrophic financial risks to the defendant. However, courts can limit the economic impact of these awards in two principal ways. First, U.S. Supreme Court decisions and state punitive damages limits can affect the magnitude of the awards. Second, courts can reduce or overturn these awards on appeal or other post-verdict procedures. This article examines the distribution and determinants of blockbuster awards and the ways in which the courts have restrained their economic impact.

Punitive damages serve a unique function in the American civil justice system. They are essentially the only civil remedy that exists to punish defendants and deter future wrongdoers as opposed to compensating plaintiffs for specific harms or preventing that harm from recurring in the future. The U.S. Supreme Court has repeatedly held that states may only authorize punitive damages for the limited purposes of punishing and deterring defendants. Because these awards exist to punish and deter, the amount that any individual plaintiff may receive does not adhere to a well-defined formula. Unlike compensatory damages which must be calibrated to a specific harm, adjudicators must decide what amount of damages will adequately achieve the more

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nebulous social goals of punishment and deterrence. When deciding what amount to award plaintiffs, adjudicators often receive very little guidance regarding the amount of punitive damages that is appropriate, leading some observers to question the extremely large awards.¹

While the U.S. Supreme Court has never instructed lower courts on exactly how they should award punitive damages, despite the potential for extremely large awards, it has provided some general guidance on what awards are unconstitutionally large. After authorizing punitive damages under the Constitution in its earlier cases,² the Court began to limit the size of awards and the process by which they are awarded in the 1990s and 2000s under the Due Process Clause of the Fourteenth Amendment. Currently, three main cases state the constitutional doctrine on punitive damages: BMW of North America, Inc. v. Gore (1996), State Farm v. Campbell (2003), and *Philip Morris USA v. Williams* (2007). All three of these cases consistently hold that the only legitimate purposes for which states can authorize punitive damages are punishment and deterrence. In all three cases, the Court seeks to limit "grossly excessive" awards that violate notions of fundamental fairness, do not provide adequate notice to parties that they may be subject to large awards, and do not further the legitimate interests of the state. In limiting these excessive awards, the Court in *BMW* provided three general guideposts that lower courts must consider when determining whether a given punitive award is "reasonable" under the Fourteenth Amendment: (1) the reprehensibility of the defendant's conduct, (2) the "disparity" between the harm inflicted on the plaintiff and the punitive damages award, and (3) the difference between the award and civil penalties authorized for similar cases.³

² Pacific Mutual Life Insurance v. Haslip, 499 U.S. 1 (1991).

¹ For a discussion of the types of guidance received by adjudicators, see Sunstein et al. (2002). In general, awarding punitive damages may be efficient from the standpoint of deterrence depending on the relationship between the probability of detecting the defendant's wrongful behavior and the amount of compensatory damages. Polinksy and Shavell (1998) provide a thorough overview of the economic theory of punitive damages, but there is little evidence that courts have adhered to the economic prescriptions for efficient levels of punitive damages.

³ These guideposts were initially announced in BMW of North America v. Gore. 517 U.S. 559 (1996).

State Farm clarified the Court's decision in *BMW* by explaining that the reprehensibility of the defendant's conduct is the most important factor in evaluating the reasonableness of a particular punitive damages award. *State Farm* will also play a central role in our empirical analysis because it provided specific guidance with respect to the appropriate magnitude of punitive damages awards. The Court in *State Farm* explicitly limited punitive damages awards under the Due Process Clause stating that "few awards exceeding a single-digit ratio between punitive and compensatory damages, to a significant degree, will satisfy due process."⁴ Using a generous interpretation of the Court's holding, no punitive damages award that exceeds the accompanying compensatory damages award by more than a factor of 10 will satisfy due process absent special circumstances.⁵ Finally, in *Williams*, the Court held that punitive damages may not be imposed to punish actions beyond the scope of the case before the court. This holding prevents juries from imposing large awards on unpopular defendants for actions they have taken outside of the scope of a given case.

After *Williams*, the Court next addressed punitive damages outside of the constitutional context in *Exxon Shipping Co. v. Baker* (2008), holding that under federal maritime law courts may not impose a punitive award greater than the accompanying compensatory award (a ratio limit of 1). While this decision is not binding outside of federal maritime law, lower courts may have inferred from *Baker* that the Supreme Court was less willing to tolerate large awards than it initially indicated in *State Farm*.

In this paper, we analyze the effect of the current punitive damages doctrine, primarily the *State Farm* decision, on blockbuster awards to determine whether it has had a significant

⁴ State Farm Mut. Auto. Ins. Co. v. Campbell, 538 U.S. 408, 415–23 (2003).

⁵ The Court has never clarified what exactly would constitute special circumstances that would justify exceeding the ratio limit imposed by *State Farm*.

negative impact on these awards. Del Rossi and Viscusi (2010) previously examined the impact of *State Farm* on an earlier sample of blockbuster awards and found that the decision was associated with a decrease in award amounts. Eisenberg and Heise (2011) and McMichael (2013) consider the effect of *State Farm* on more typical cases in the Civil Justice Survey of State Courts data and, in contrast to the results for blockbuster cases, find that it has not had a significant negative effect on punitive damages awards in general.⁶

When examining blockbuster awards in this paper, we consider a longer time frame post-*State Farm* than did Del Rossi and Viscusi (2010) and examine the effect of *State Farm* on the amount of punitive damages awarded in individual cases as well as on the probability that a given award involved a punitive to compensatory damages ratio of greater than 10—the limit imposed by *State Farm*. We find that there is a negative influence of *State Farm* on award size and the probability of exceeding a single digit ratio for blockbuster awards.

Additionally, we explore the distribution of blockbuster punitive damages and find that it is similar to the distribution of natural disasters and other catastrophic events. In particular, blockbuster awards follow a fat-tailed distribution, meaning that extremely large awards (catastrophic events) occur much more often than if awards were distributed normally. Thus, awards following a fat-tailed distribution are more difficult to predict than normally distributed awards. The nature of this distribution has important legal implications, as the U.S. Supreme Court has based much of its case law on the constitutional requirement that punitive damages

⁶ Eisenberg and Heise (2011) consider whether juries award more punitive damages than judges. As part of this analysis, they find that of the four years for which they have data (1992, 1996, 2001, and 2005) on punitive damages across the country, the difference between how judge and juries award punitive damages was statistically significant only in 2005. They consider *State Farm* as a possible reason for this difference, but conclude that the shift in awarding behavior is not attributable to *State Farm*. McMichael (2013) conducts explicit tests for an effect of State Farm on punitive damages awards similar to those used here and finds no evidence of a significant negative effect and some evidence of a significant positive effect of State Farm on punitive damages award amounts.

remain predictable so that defendants have adequate notice of their potential liability.⁷ We show that while blockbuster punitive damages follow a fat-tailed distribution both before and after the Court's decision in *State Farm*, the distribution shifted to the left after *State Farm*, i.e., the tail is not as fat, which is consistent with a restraining effect on blockbuster awards. In other words, after *State Farm*, defendants should be better able to predict extremely large punitive damages awards.

II. FAT-TAILED DISTRIBUTIONS AND PUNITIVE DAMAGES

Blockbuster punitive damages awards by their very nature are unusual outliers among all punitive damages awards. However, even within the category of awards that qualify as blockbuster awards, the distribution includes very extreme outliers relative to other blockbuster awards. The extreme values of blockbuster punitive damages awards at the upper tail of the blockbuster award distribution exhibit many parallels with the distribution of losses from natural disasters and other catastrophic events. The losses from natural disasters are not normally distributed but tend to follow fat-tailed distributions.⁸ With these distributions, the largest loss may be several times larger than the next largest loss. For example, the 2004 Sumatra earthquake which resulted in the deaths of 227,898 people is the second worst earthquake on record. The worst earthquake on record, which occurred in 1976 in Tangshan, China, killed 655,000 people according to many estimates—almost three times as many as the 2004 event.⁹ Disasters resulting from human error also tend to follow fat-tailed distributions. For example, the widely publicized

⁷ See BMW, 517 U.S. at 560.

⁸ See U.S. Geological Service Fact Sheet, "Natural Disasters – Forecasting Economic and Life Losses," http://pubs.usgs.gov/fs/natural-disasters/index.html.

⁹ See U.S. Geological Survey, Largest Earthquakes in the World Since 1900, http://earthquake.usgs.gov/earthquakes/world/10_largest_world.php.

Deepwater Horizon disaster in 2010 caused the spill of 4.9 million barrels of oil. However, the worst oil disaster on record occurred in 1910 and resulted in the spill of 9 million barrels.¹⁰ Viscusi and Zeckhauser (2011) discuss these and other examples of large-scale losses that follow fat-tailed distributions.

In addition to natural and human-made disasters, other catastrophic events may follow fat-tailed distributions. For example, Bremmer and Keat (2009) demonstrate that volatile political events, which can have catastrophic effects on domestic and international markets, can follow a fat-tailed distribution. They note that because catastrophic political events like the confrontation between Georgia and Russia in 2008 and the 1998 Russian financial meltdown follow a fat-tailed distribution, they happen more frequently than investors—who are usually adept at handling risk—predict. They explain that fat tails "represent the risk that a particular event will occur that appears so catastrophically damaging, unlikely to happen, and difficult to predict, that many of us choose to simply ignore it" (Bremmer and Keat 2009, pp. 2 - 3). While ignoring extreme or catastrophic events may be justified when they are distributed normally, doing so when they follow a fat-tailed distribution can prove problematic.

Because normal distributions do not capture the fat-tailed behavior of disasters and other catastrophic events, analyses of losses from natural disasters have employed power-law distributions, which have fat tails, to describe these losses.¹¹ With this distribution, the event with the greatest impact may be several times as large as the next most significant event, which is a type of variation not exhibited by events that must adhere to a normal distribution. The United States Geological Survey generated Figure 1 to demonstrate the power-law distributions

¹⁰ See World's Largest Oil Spills Map, Geology, http://geology.com/articles/largest-oil-spills-map/

¹¹ A variable p(x) has a power law distribution if it can be characterized as $p(x) = cx^{\beta}$ where c and d are constants so that $\log(p(x)) = \log(c) + d\log(x)$.

of various types of natural disasters. Both axes are on a logarithmic scale so that they progress in orders of magnitude. In this graph, the measure of loss shown on the horizontal axis is the number of deaths, and the vertical axis measures the cumulative number of each type of event per year. Each set of points illustrates the power-law distribution of a different type of natural disaster including hurricanes, tornadoes, and earthquakes.

Blockbuster punitive damages awards likewise follow a fat-tailed distribution similar to the distribution of various natural disasters. While the cutoff for blockbuster status is a nontrivial \$100 million, many of these awards exceed this cutoff by one, two, or even three orders of magnitude. Instead of a normal distribution, or even a lognormal distribution, a power law distribution best describes blockbuster awards.

Figure 2 illustrates the power law distribution that blockbuster punitive damages awards follow. It reports the "inverse cumulative distribution function" of 130 blockbuster awards that we identified between 1980 and 2012 and which we analyze below.¹² The "cumulative distribution function" in Figure 2 is "inverse" in the sense that each point represents the size of the award (plotted along the horizontal axis) and the proportion of awards that are smaller (plotted along the vertical axis) as opposed to the proportion of larger awards in a true cumulative distribution function. Figure 2 also plots a power trend line through the data to show how well this data is approximated by a power function.¹³ Both the vertical and horizontal axes

¹² Where a true cumulative distribution function of random variable X is given by $F(x) = P(X \le x)$, the inverse cumulative distribution function presented in Figure 2 is given by F(x) = P(X > x).

¹³ We exclude the two largest cases (*Engle* and *Middleton*, which are discussed at length below) from this figure. We exclude these cases only because doing so is necessary to construct this and later inverse cumulative distribution functions. Because the inverse cumulative distribution function is given by F(x) = P(X > x), the largest award will, by definition, have a zero probability of taking on a value larger than *x*. This zero probability is not problematic in general, but because we use logarithmic scales for both axes, we must either assign a positive probability for the largest award or remove it from the analysis. Because any probability we assign would be arbitrary, we remove the largest award (which was imposed in *Engle*) from the analysis. We also remove the second largest award (which was imposed in *Engle*) from the analysis because we later include figures that group awards by whether they were imposed before or after *State Farm* and *Middleton* is excluded as the largest award in the after-*State Farm*

use a logarithmic scale so that they progress in orders of magnitude. The fact that the data are essentially linear on this logarithmic scale demonstrates the fat-tailed aspect of the distribution of blockbuster awards.

Similar to the largest losses from natural disasters, the largest punitive awards may be several times larger than the next largest award. For example, consider the two largest awards in Figure 2. The largest award is \$35 billion while the second largest is \$14 billion—less than half the size of the largest. Including the two largest awards (*Engle v. R.J. Reynolds Tobacco* and *Middleton v. Collins*), exacerbates the fat tail even more. These two awards are four times (*Middleton*) and five times (Engle) larger than the largest award presented in Figure 2.

With such large awards possible, defendants will likely not be able to predict what amount of monetary liability they may face for their actions. State policymakers and both state and federal courts have recognized that defendants may lack adequate notice of the liability they may face and have stepped in to afford a number of protections against excessive punitive damages. Many states have enacted specific tort reforms limiting the total amount of punitive damages available to plaintiffs and/or the ratio between the amount of compensatory and punitive damages.¹⁴ For example, Florida currently limits punitive damages to the higher of \$500,000 or three times the amount of compensatory damages with these limits increasing if the defendant knew his actions risked injury to or intentionally injured the plaintiff. Some states, e.g., Louisiana, only authorize punitive damages for certain types of cases. Nebraska has simply banned punitive damages.

group. We exclude it here as well to maintain comparability between figures. However, including it here does not change the conclusion about the fat-tailed nature of the distribution of blockbuster awards.

¹⁴ Avraham (2011) provides a comprehensive list of state tort reforms including those directed at punitive damages.

III. THE EFFECT OF THE CURRENT CONSTITUTIONAL DOCTRINE ON PUNITIVE DAMAGES

As in previous blockbuster award studies, we identified blockbuster awards by conducting a broad and thorough search of a variety of sources including, but not limited to: Westlaw databases, Lexis-Nexis databases, LEXIS combined jury verdicts and settlements, and VerdictSearch's annual "Top 100 Jury Verdicts" for 2009, 2010, 2011, and 2012. We obtained basic case information, including the year of the award, the state where the case was heard, the type of industry involved, and the amounts of punitive and compensatory damages. We obtained this information primarily from news articles and, where possible, court documents. Our search yielded a total of 32 new blockbuster awards which, combined with previously identified awards in Del Rossi and Viscusi (2010), brings the total to 132 blockbuster awards. Of these 32 new awards, eight were imposed between 1981 and 2008 and were not identified in previous work. Twenty-four of the newly identified awards occurred between 2009 and 2012, which was beyond the time frame considered in earlier studies.

Before delving into the analysis of the effect of the *State Farm* decision on punitive damages awards, we report in Table 1 and Figure 3 a basic overview of the distribution of blockbuster punitive damages awards. Table 1 organizes the 132 awards we have identified into deciles—with the first group having two more awards than the others—and reports basic statistics for each. The increase in means between the lower deciles and the higher deciles illustrates how quickly blockbuster awards increase at the higher end of the distribution. For example, the increase (in \$millions) in the mean between deciles is never more than 50 below the sixth decile while the increase is 154 between the seventh and eighth, 476 between the eighth and ninth, and over 30,000 between the ninth and tenth. This huge jump in the final decile

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reflects the fat-tailed nature of the distribution. The ranges covered by the deciles tell a similar story.

Figure 3 illustrates the right skew of blockbuster punitive damages awards. Even after taking the logarithm of award amounts, this skew is obvious and displays a fatter tail than would a lognormal distribution. The skewness for award levels is 7.9 while the skewness for the logarithms of awards is 2.3. These measures of skewness along with the figures demonstrate the positive skew of the distribution of blockbuster awards, meaning that the right tail of the distribution (which contains the large, outlying awards) is relatively longer than the left tail of the distribution. In terms of comparison, the skewness of a standard normal distribution is zero indicating that the left and right tails of the distribution are essentially the same.

Figure 3 also illustrates the fat-tailed nature of blockbuster awards fairly well, as several of the intervals at the higher end of the distribution contain no awards demonstrating that the largest awards are much larger than the next largest awards. Figure 2 above also clearly demonstrates the fat tail of blockbuster punitive damages awards. To more accurately measure the thickness of the tails of the distribution of blockbuster punitive damages, we calculate excess kurtosis, which measures how heavy the tails of a distribution are. For the distribution of blockbuster punitive damages, the excess kurtosis is 63.5, and for the distribution of the natural logarithm of blockbuster punitive damages, the excess kurtosis is 7.2. Both of these measures indicate that the distribution has fat tails. The typical point of comparison for excess kurtosis is the standard normal which has an excess kurtosis of zero. An excess kurtosis greater than zero indicates that a distribution has thick tails, so the distribution of blockbuster punitive damages clearly has thick tails. In terms of comparison to other common distributions, a logistic

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distribution with a zero mean, no skew, and a variance of 1 has an excess kurtosis of 1.2, and a uniform distribution with a zero mean and a variance of 1 has an excess kurtosis of -1.2.

Figure 4 provides an overview of the annual total amount of blockbuster punitive damages that have been awarded. These awards were somewhat sporadic through the 1980s, producing a saw-tooth pattern of total punitive awards. In the 1990s, the total amount of punitive damages awarded as result of blockbuster cases increased. Two important changes to the total amount of punitive damages awarded in terms of blockbuster awards occurred in the early 2000s. First, the total blockbuster award amount per year began to decline. Second, that decline occurred without much volatility. We analyze whether this somewhat dramatic decline was the result of the Supreme Court's decision in *State Farm*.

To begin that analysis, Figure 5 presents the "inverse" cumulative distribution function of blockbuster punitive damages awards before and after the *State Farm* decision. Both before and after *State Farm*, blockbuster awards follow a power law distribution. However, the distribution of awards shifted downward after the *State Farm* decision. After that decision, the awards are generally smaller. The pre-*State Farm* awards clearly follow a distribution with a fatter tail than that of the post-*State Farm* awards suggesting that *State Farm* may have had a greater effect on the largest punitive damages awards. Considering only the four largest cases (*Engle*,¹⁵ *Middleton v. Collins*,¹⁶ *Bullock v. Philip Morris, Inc.*,¹⁷ and *Exxon Mobil Corporation v. Alabama Dept. of*

¹⁵ Engle v. Liggett Group, Inc., 945 So. 2d 1246 (Fla. 2006). This citation is for the most recent version of this case, not the case in which the jury awarded the blockbuster amount. A jury originally imposed a blockbuster award in 2000. See Section IV below for a more detailed discussion of Engle.

¹⁶ This case has not been reported in an official reporter. The blockbuster award was originally imposed in 2011. See Top 100 Verdicts of 2011.

¹⁷ *Bullock v. Philip Morris USA, Inc.*, 198 Cal. App. 4th 543 (2011), review denied (Nov. 30, 2011). This citation is for the most recent version of this case, not the case in which the jury awarded the blockbuster amount. A jury originally imposed a blockbuster award in 2002. See Section IV below for a more detailed discussion of Bullock.

*Conservation and Natural Resources*¹⁸), the greater number of outlying awards prior to *State Farm* (three of the four occurred prior to State Farm) suggests that lower courts became less willing to hand down extremely large judgments after *State Farm*. We discuss these four cases in greater detail below.

Turning to our regression analysis, we consider two primary questions. First, have individual award sizes decreased in the post-*State Farm* era, and second, has the single-digit ratio cap imposed by *State Farm* reduced the number of blockbuster cases with high (greater than 10) punitive to compensatory damages ratios? For both, we use the following general specification for the determinants of different measures, *Y*_{*it*}, of punitive damages, which is an approach consistent with Del Rossi and Viscusi (2010):

 $Y_{it} = \propto + \beta \log(Compensatory \ Damages)_i + \delta \log(Compensatory \ Damages)_i * State \ Farm_{it} + \gamma Litigant \ Type_i + \theta Industry_i + \varphi State_i + \omega Bench \ Trial_i + \varepsilon_{it}.$

In this equation, *i* indexes individual cases and *t* indexes time.

For the first question of the effect of *State Farm* on award size, *Y*_{it} is the natural logarithm of punitive damages. For the second question concerning the effect of *State Farm* on the probability that an individual award violates the single digit ratio, *Y*_{it} is an indicator for whether the punitive to compensatory damages ratio exceeds 10. The explanatory variables include the natural logarithm of compensatory damages, and a higher level of compensatory is associated with a higher level of punitive damages in general.¹⁹ After *State Farm*, higher levels of compensatory damages, but an increase in compensatory damages should be associated with a smaller increase in

¹⁸ See *Exxon Mobil Corp. v. Alabama Dept. of Conservation & Natural Res.*, 986 So. 2d 1093, 1097-98 (Ala. 2007). The blockbuster award in this case was originally imposed in 2003.

¹⁹ See Eisenberg et al. (2002), Hersch and Viscusi (2004), Del Rossi and Viscusi (2010), and Eisenberg and Heise (2011).

punitive damages after *State Farm* than before. Because *State Farm* imposed a cap on the ratio between punitive and compensatory damages, the average and median ratios should decrease after that decision, which means that an additional dollar of compensatory damages should be, in general, worth fewer additional dollars of punitive damages after *State Farm*.

State Farm is an indicator for whether the case was heard after 2003 and is a principal variable of interest. This variable takes a value of 1 for every case heard in 2004 and thereafter. While the Supreme Court decided the *State Farm* case in April 2003, the *State Farm* indicator does not take a value of 1 until 2004 to allow for the decision to be fully implemented by lower courts. We interact the *State Farm* indicator with the log of compensatory damages to determine if the relationship between compensatory and punitive damages changed after the *State Farm* decision, as one would predict based on the decision's guidance with respect to appropriate ratios of punitive damages to compensatory damages.

Litigant Type consists of a single indicator variable that controls for business and individual litigants, and the excluded category is other litigant types (governments, etc.). *Industry* is a vector of indicator variables for different industry types including the automobile, tobacco, finance/investment/insurance, energy/chemical, and pharmaceutical/health industries. This vector also includes an indicator for whether a case involved a violent crime. *State* is a vector including indicator variables for whether a case was decided in Texas or California—two states associated with large punitive awards. Finally, *Bench Trial* takes a value of 1 when a case was decided by a judge, not a jury.

In analyzing the effect of *State Farm*, we exclude several cases. First, we exclude *Garamendi v. Altus Finance, S.A.* and *Chopourian v. Catholic Healthcare West* from all regressions because the courts in these cases awarded no compensatory damages so that the ratio

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of punitive damages to compensatory damages for these cases is not defined. Second, we exclude both *Middleton* and *Engle*—the two largest cases—from some regressions. These two cases, as well as *Bullock* and *Exxon Mobil* are discussed in greater detail in Section IV below.

Table 2 reports two sets of OLS regression results for the log of punitive damages. The first column reports a simple regression with only two regressors relating to the relationship between compensatory and punitive damages. The results indicate that this relationship changed after the Supreme Court decided *State Farm*. Because both the dependent and independent variables are in log form, the coefficient on the log of compensatory damages can be interpreted as an elasticity. Prior to *State Farm*, a 1% change in compensatory damages is associated with a 0.172% increase in punitive damages. The coefficient on the interaction between the log of compensatory damages and the *State Farm* indicator represents how that elasticity changed after the *State Farm* decision. The elasticity decreased by 0.036 after *State Farm*, leading to a net elasticity of about 0.14.

The second column of estimates in Table 2 repeats the same analysis adding all of the control variables discussed as part of the main specification. Adding these controls does not alter the statistical significance of the results for *State Farm*; however, the magnitudes do change slightly. Before *State Farm* the elasticity between compensatory and punitive damages is about 0.20 and after, it decreases to about 0.17. Our estimated effect of *State Farm* is smaller than the effect found by Del Rossi and Viscusi (2010). The difference in the results indicates that the effect of *State Farm* has weakened over time, as we include four more years of blockbuster cases than Del Rossi and Viscusi (2010).²⁰ However, despite the inclusion of four years of cases after

²⁰ We also include eight additional cases prior to *State Farm* that were not included in earlier work. However, dropping those eight cases from our analysis does not meaningfully alter the regression results. The coefficient on the logarithm of compensatory damages increases to 0.220 while the coefficient on the interaction term between

State Farm, our estimated net elasticity between compensatory and punitive damages of 0.17 is similar to Del Rossi and Viscusi's (2010) estimated net elasticity of 0.13. This suggests that the relationship between compensatory and punitive damages is relatively stable.

We repeated the regressions reported in Table 2 with an indicator variable for whether a case was decided after *Exxon Shipping Co. v. Baker*. In that case the U.S. Supreme Court limited the punitive to compensatory damages ratio to 1 under federal maritime law. While that decision does not apply to cases generally, Justice Ginsburg indicated in her opinion that the Court favored stricter limitations on punitive damages. Therefore, we included this case in our analysis to examine whether *State Farm* is the driving force behind any observed change in the elasticity between punitive and compensatory damages. The coefficient on the interaction between the indicator for the *Exxon Shipping* case and the log of compensatory damages was never statistically significant, and none of the other results changed in any meaningful way.

Because blockbuster awards follow a power law distribution, as opposed to a normal or lognormal distribution, we also estimate quantile regression models, which are better suited to analyzing the effects of *State Farm* across the distribution of awards. The number of variables that can be used in a quantile regression is limited, especially with a small sample size, because these regressions concentrate on the determinants of award amounts for specific quantiles accounting for the probability of being in a given quantile. Therefore, we do not include all of the control variables we use in the OLS regressions. Table 3 reports a series of quantile regressions are robust to large outliers, we do not exclude any large awards from this analysis; however, because the award in the *Engle* cigarette class action case dwarfed all other awards up to that time, all of the

compensatory damages and the State Farm indicator increases to a -0.034. Therefore the inclusion of these eight cases excluded by Del Rossi and Viscusi (2010) is not responsible for the weakening of the *State Farm* effect.

regressions in Table 3 control for the effect *Engle* may have had on the relationship between compensatory and punitive damages. Consistent with the OLS regression results, the *State Farm* decision has a significant negative effect on the relationship between compensatory and punitive damages for each of the quantiles. The *State Farm* interaction term is significant, implying *State Farm* changed the relationship between compensatory and punitive damages so that an additional dollar of compensatory damages is associated with a smaller increase in punitive damages after the Supreme Court handed down the decision.

The pattern of the coefficients for the *State Farm* interaction term is especially interesting. Looking at the magnitudes of the *State Farm* interaction terms across the three sets of estimates, the *State Farm* decision influenced awards at the higher end of the distribution more substantially than those at the lower end. The coefficient on the interaction term has similar values of -0.031 for the 25th percentile and -0.033 for the median. However, the coefficient more than doubles in absolute value to -0.069 for the 75th percentile, indicating that *State Farm* has had the biggest impact on the most extreme awards.

The combined evidence from the OLS and quantile regressions demonstrates that the *State Farm* decision changed the relationship between compensatory and punitive damages in blockbuster cases. The evidence from the regressions is consistent with Figure 5 above, which plots power function trend lines for awards imposed before and after the *State Farm* decision. These trend lines in Figure 5 demonstrate visually that while blockbuster punitive damages awards still follow a power law distribution after *State Farm*, award levels have been lower since 2004. Consistent with the results of the quantile regressions, Figure 5 also demonstrates that the effect of *State Farm* has been strongest for the larger awards, as the power law trend line for the awards before *State Farm* decreases more quickly than the trend line for the awards before *State*

Farm. Overall, *State Farm* shifted the inverse cumulative distribution function of blockbuster cases down.²¹ Similar to other figures and tables, the trend lines plotted in Figure 5 exclude the two largest cases, while the quantile regressions do not.

To more accurately capture the evidence demonstrated visually by Figure 5, we calculate the skewness and excess kurtosis of the distributions of blockbuster punitive damages awards before and after *State Farm* was decided. The distribution of pre-*State Farm* blockbuster awards has an excess kurtosis of 70.6, and this decreases to 55 for the distribution of post-*State Farm* awards, indicating that the post-*State Farm* distribution has a thinner tail than the pre-*State Farm* distribution. The skewness of the distributions of pre- and post-State Farm award amounts follows a similar pattern. The distribution of pre-*State Farm* awards has a skewness of 8.2, and this decreases to 7.4 for the distribution of post-*State Farm* awards, indicating that the post-*State Farm* awards.

Excluding the largest case from both the pre- and post-*State Farm* distributions (*Engle* and *Middleton* respectively) makes the differences between the two distributions even more dramatic. The excess kurtosis of the pre-*State Farm* distribution is 44.6, and this decreases to 10 for the post-State Farm distribution. The skewness of the pre-*State Farm* distribution is 6.3 while the skewness of the post-*State Farm* distribution is 2.8. These measures of excess kurtosis and skewness indicate that while the distributions of blockbuster awards before and after *State Farm* have thick tails and positive skews, the tail becomes thinner and the skew less dramatic after

²¹ To confirm that the shift observed in the inverse cumulative distribution function represents a statistically significant change in the power law function that describes the cases before and after State Farm, we ranked the cases before and after state farm so that the following power law function describes them: *award rank* = $(award size)^{\beta}$. The exponent β that characterizes the power law relationship is significantly different for awards imposed before and after state farm (t-statistic = 22.67).

State Farm. This evidence is consistent with the results of our regression analysis indicating that State Farm is associated with a decrease in blockbuster award amounts.

The evidence from the regressions and the visual evidence from Figure 5 indicate that *State Farm* has significantly influenced how lower courts impose punitive damages. However, this evidence elucidates *State Farm*'s effect on blockbuster punitive damages awards generally. Because *State Farm* imposed a single digit cap on the ratio between punitive and compensatory damages, it is important to determine if the observed *State Farm* effect influences punitive damages awards in general or is concentrated around the single digit threshold. Therefore, we conclude our empirical analysis of the effect *State Farm* by directly testing whether lower courts are more likely to impose awards with single digit punitive to compensatory damages ratios after *State Farm*.

Prior to the *State Farm* decision, federal and state courts imposed a total of 77 blockbuster punitive damages awards. Of these awards, 43 (or 56%) involved a double-digit ratio of punitive to compensatory damages. Since the U.S. Supreme Court decided *State Farm*, federal and state courts have imposed a total of 55 blockbuster awards.²² Of these awards, 15 (or 28%) involved a double-digit ratio of punitive to compensatory damages. The decrease in the percentage of blockbuster awards exceeding the single digit ratio imposed by State Farm suggests that State Farm may have changed how lower courts impose punitive damages awards, and we use probit models to examine this potential change more closely.

Table 4 reports the results from two probit regressions in which the dependent variable is an indicator for whether the punitive to compensatory damages ratio in a case exceeded $10.^{23}$ All

²² Two of these awards involved no accompanying compensatory damages award. We drop these cases in our analysis of ratio of punitive to compensatory damages.

²³ The Court in *State Farm* actually said that "few awards exceeding a single-digit ratio between punitive and compensatory damages, to a significant degree, will satisfy due process." 538 U.S. at 425. However, we use an

coefficients have been transformed to correspond to marginal effects. In both the simple regression and the regression with additional controls, the coefficient on the log of compensatory damages is negative, indicating that for higher levels of compensatory damages, the total award in a case is less likely to violate the single digit ratio cap. The coefficient on the *State Farm* interaction term is negative and significant in both specifications, so the *State Farm* decision changed how courts award punitive damages in relation to compensatory damages. For a given level of compensatory damages, after the Court decided *State Farm*, lower courts became less likely to impose an award in which the punitive to compensatory damages ratio exceeded 10. This effect is relatively small, however, as it reduces the probability of exceeding the single digit ratio upper bound by only 0.02. The effect of *State Farm* is consequently manifested more broadly throughout the blockbuster award distribution and is not narrowly concentrated around the value of the suggested cap.

IV. THE FOUR LARGEST BLOCKBUSTER PUNITIVE DAMAGES CASES

Our analysis indicates that *State Farm* had a broader effect than simply decreasing the probability of crossing the double-digit threshold. However, that analysis has considered only the award imposed on the defendant and not what the defendant ultimately paid to the plaintiff. In practice, defendants may not pay the full award amount for a variety of reasons including post-verdict settlements, reduction of the award through remittitur, or reduction of the award by a higher court. In this section, we discuss the four largest blockbuster punitive damages awards

indicator for whether a given ratio exceeds 10 because clever attorneys may point out that 9.999... = 10 and that 9.999... is still a single digit. This formulation of the indicator variable allows for the widest possible interpretation of the Supreme Court's ruling. Only two cases involved a ratio of exactly 10.

ever imposed to illustrate the post-verdict litigation processes that may reduce the amount the defendant must ultimately pay in a given case.

Viscusi (2004) discusses why defendants may not ultimately pay the full award amount and provides additional information on the litigation history of 64 blockbuster awards. Of those 64 awards, ten were reversed on appeal while only one was affirmed on appeal. One case required a retrial while three awards were reduced on appeal, and three more settled during the appellate process. One award was reduced by the trial court, and 21 cases settled prior to the appellate process. In two cases, the defendant did not pay the punitive damages award.

The discussion here only concerns four cases, but it provides much greater detail on the litigation parties engage in after a court imposes a blockbuster punitive damages award. Unlike earlier work which only provided the post-award outcome of a case, this discussion illuminates the litigation histories of the four largest cases to illustrate not only that defendants may not ultimately pay the initial blockbuster award but also the complexity and length of the litigation involved in these types of cases.

State courts imposed all four of the most extreme awards, and three were imposed prior to *State Farm* while one was awarded after. Only one of these awards survived subsequent litigation, but it is highly unlikely that even in that case the defendant will pay the full award. Because each case was quite complex—one case involved over 57,000 pages of litigation transcripts—the discussion here is necessarily a succinct overview.

A. Engle v. R.J. Reynolds Tobacco

Beginning with what may be the most widely discussed punitive damages case, a Florida circuit court decided *Engle v. R.J. Reynolds Tobacco* in 2000 (prior to *State Farm*), awarding \$144.8 billion in punitive damages. This award remains the largest (after adjusting for inflation)

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punitive damages award in history. Coupled with an initial compensatory damages award of \$12.7 million, this case involved a punitive to compensatory damages ratio of over 11,000.²⁴ Here we explore the process that generated this award as well as why this award was never paid. The case began in 1994 as a class action lawsuit on behalf of all U.S. citizens injured by addiction to cigarettes. The plaintiffs alleged that the tobacco industry intentionally manipulated the nicotine in cigarettes to make them more addictive and concealed information concerning these actions and the addictive nature of nicotine in cigarettes. A Florida appellate court reduced the class certification to only Floridians harmed by cigarettes and held that while liability could be determined on a class basis, damages should be determined individually. After the class was certified, the trial court conceived a trial plan involving three phases to handle a class of over 500,000 individuals.

In phase I, a jury of six found R.J. Reynolds and other tobacco companies liable for a number of alleged harms (many of which involved some sort of fraud) and determined that punitive damages were allowable without setting a specific amount of either compensatory or punitive damages. Phase II involved the determination of compensatory damages for two representative class members and the assessment of a lump sum of punitive damages for the entire class. When the trial court announced it would impose punitive damages as a lump sum prior to determining individual damages for each class member, the tobacco companies appealed the use of this procedure, but the Florida Supreme Court allowed the trial court to use this procedure while allowing the tobacco companies to appeal this issue again in future litigation. Imposing punitive damages as a lump sum for the entire class, the jury arrived at a total award of

²⁴ This ratio is somewhat misleading, as the compensatory damages award was for a few representative class members while the punitive damages award was for the entire class.

\$144.8 billion.²⁵ The size of the punitive damages award and the novel procedures which led to its imposition prompted action by the state legislature and state attorney general. The legislature passed a law preventing large punitive damages awards (such as the one in *Engle*) from being paid out while a case proceeded through the appeals process. The attorney general entered the appeals process on the side of the tobacco companies arguing that the procedure used by the trial court to impose the large punitive award violated state law and that the size of the award violated the Due Process Clause of the Fourteenth Amendment.

The Florida District Court of Appeal issued an opinion in 2003 that was favorable to the tobacco companies.²⁶ The court held that the class of smokers in the case was not properly certified and struck down the punitive damages award on several grounds. First, the court held that awarding punitive damages for the entire class without additional evidence being presented was improper. Second, even if the punitive damages award had been properly determined, it was excessive under both state and federal law. The court explained that a punitive award could not be so large as to bankrupt/destroy a defendant under Florida law and, citing *State Farm*, held that such a large award violated the Due Process Clause. Finally, the court explained that the trial court had erred in not considering the Florida Settlement Agreement, which assessed punitive damages against tobacco companies on behalf of the citizens of Florida, when setting the punitive damages award in *Engle*.

The Supreme Court of Florida held that the Florida Settlement Agreement did not prohibit the trial court from awarding punitive damages, but it did uphold the appellate court's ruling overturning the punitive damages award. The court explained that while a lower court

²⁵ Individual class members were supposed to receive an equal portion of the lump punitive damages award regardless of their individual compensatory award.

²⁶ Liggett Group Inc. v. Engle, 853 So. 2d 434 (Fla. Dist. Ct. App. 2003) decision approved in part, quashed in part sub nom. Engle v. Liggett Group, Inc., 945 So. 2d 1246 (Fla. 2006)

need not award a specific amount of compensatory damages prior to determining whether the plaintiff is entitled to punitive damages, it must fully consider all aspects of liability prior to determining that punitive damages are allowable. In this case, the jury did not do so. Additionally, the court held that prior to setting a specific amount of punitive damages, lower courts must set a specific amount of compensatory damages because it must consider whether the whether the punitive award is reasonably related to the harm caused (as reflected by the total compensatory award). Here, the jury did not know the total compensatory award before determining the total punitive award which violated both state and federal law. The court cited *State Farm* in striking down the lower court's procedure of determining punitive damages prior to knowing the total harm caused and held for the first time under Florida law that the punitive award must bear a reasonable relationship to the compensatory award. Ultimately, the court was more friendly toward the plaintiffs than the appellate court but upheld the ruling striking down the punitive award.

While the Supreme Court of Florida affirmed the ruling striking down the blockbuster award, the litigation surrounding *Engle* continued. A full review of the *Engle* "progeny cases" is well beyond the scope of this work, but Henderson and Twerski (2010) provide a thorough discussion of tobacco litigation including the cases spawned by *Engle*.

B. Bullock v. Philip Morris USA

Unlike the award in *Engle*, the punitive damages award in *Bullock v. Philip Morris USA* survived the appeals process, albeit at a lower level than initially awarded.²⁷ In this case, a single plaintiff filed a suit against Philip Morris for lost earnings and medical expenses stemming from

²⁷ This case changed names throughout the litigation. Bullock v. Morris, BC249171, 2002 WL 31833905 (Cal. App. Dep't Super. Ct. Dec. 18, 2002) aff'd sub nom. Bullock v. Philip Morris USA, Inc., 138 Cal. App. 4th 1029, 42 Cal. Rptr. 3d 140 (2006) review granted and opinion superseded, 141 P.3d 718 (Cal. 2006)

the plaintiff's smoking. In 2002 (prior to *State Farm*), the jury awarded \$850,000 in compensatory damages and \$28 billion in punitive damages, resulting in a punitive to compensatory damages ratio of nearly 33,000. After the defendant challenged the punitive award, the court, relying on the U.S. Supreme Court's decision in *BMW v. Gore*, determined that the punitive award was legally excessive. The court offered the plaintiff a choice: accept either a new trial or a reduction of the punitive damages award to \$28 million, which reduced the punitive award by a factor of 1,000. The plaintiff accepted the reduced award, and both the plaintiff and defendant appealed the trial court's ruling.

The appellate court rejected Philip Morris's arguments that the reduced award was unreasonable under *State Farm* (which was decided between the trial and appellate proceedings in this litigation) and upheld the \$28 million award. Philip Morris appealed to the California Supreme Court but that court deferred hearing the case until the Supreme Court of the United States decided *Philip Morris USA v. Williams*,²⁸ which was expected to provide guidance on some of the issues in the *Bullock* litigation. After the Court decided *Williams*, holding that a punitive damages award could not punish the defendant for harms to individuals outside the scope of the current litigation, the California Supreme Court for reconsideration. There, the court of appeal rejected Philip Morris's indirect challenges to the punitive award (e.g. evidentiary challenges) and its direct challenges to the trial court's findings. However, the court ordered a new trial on the issue of punitive damages because the jury instructions used in the original trial were not consistent with the Supreme Court's holding in *Williams*. The jury instructions did not adequately explain to the jury that any

²⁸ Philip Morris USA v. Williams, 549 U.S. 346 (2007)

punitive award could not punish the defendant for harms inflicted on parties outside the litigation.

At the new trial, the jury awarded the plaintiffs \$13.8 million in punitive damages. Philip Morris appealed this award, but the court of appeal upheld the award and rejected Philip Morris's arguments after a thorough analysis of the reasonability of a punitive damages award under *State Farm* and state court decisions. Interestingly, the court explained that courts have some discretion in applying the *State Farm* ratio cap and that the three guideposts announced in *BMW* and reiterated in *State Farm* ultimately guide the reasonability analysis of a given award.

C. Exxon Mobil Corp. v. Alabama Department of Conservation and Natural Resources

Unlike the previous two cases, *Exxon Mobil Corp. v. Alabama Department of Conservation and Natural Resources* (DCNR) does not involve the tobacco industry.²⁹ This case began when Exxon leased land from the state of Alabama to develop natural gas reserves under Mobile Bay.³⁰ After obtaining the leases, Exxon was supposed to pay Alabama royalties on the gas it extracted. Prior to the discovery of reserves under Mobile Bay, Alabama had used a standard lease for a number of years, but because of the value of the new reserves, the state rewrote the lease so that it better maximized the royalties owed to the state and was more favorable to the state in general. In the early 1980s, Exxon paid over \$550 million in nonrefundable bonuses to obtain leases for the reserves and agreed to pay royalties under the terms of the newly written leases. In 1993, the state discovered that Exxon was using certain practices when calculating royalties that the state believed were inconsistent with the procedures established in the leases. Exxon paid an additional \$100 million in royalties to the state, but the

²⁹ Exxon Mobil Corp. v. Alabama Dept. of Conservation & Natural Res., 986 So. 2d 1093, 1097-98 (Ala. 2007)

³⁰ Exxon changed its name to Exxon Mobile in the course of this litigation, but we refer to that company as Exxon for conciseness.

state demanded \$50 million which Exxon refused to pay. Exxon filed a suit seeking a declaratory judgment that it did not owe the additional amount and the state counterclaimed under theories of fraud and breach of contract seeking the additional amount and punitive damages.

At trial, the jury awarded the state over \$60 million in additional royalties and \$27 million in interest. Additionally, the jury imposed a punitive damages award of \$3.42 billion. As required by Alabama law, the court held a post-trial hearing on the amount of punitive damages but determined not to reduce the award. Exxon appealed the decision, and the appellate court ordered a new trial because the trial court admitted inadmissible evidence. A retrial took place in October 2003 (*State Farm* was decided in April 2003), and the jury awarded the state \$63 million in additional royalties, over \$39 million in interest, and \$11.8 billion in punitive damages.³¹ This verdict resulted in a punitive to compensatory damages ratio of approximately 187. At a post-trial hearing, the court reduced the punitive damages award to \$3.5 billion but refused to grant a new trial.

Exxon appealed the new award, and the Supreme Court of Alabama detailed exactly what was required when paying royalties under the lease. The court upheld the lower court's judgments on the state's breach of contract claims but reversed the lower court on the fraud claims (the punitive damages in this case were based solely on the fraud claims). Under Alabama law, punitive damages could not be awarded for breach of contract claims, so the Supreme Court of Alabama overturned the punitive damages award and reduced the compensatory damages amount by the amount that was awarded for the fraud claims so that the final award was about \$51 million in total. Like *Engle*, this case involved a blockbuster punitive damages award, but after proceeding through the appeals process, that award ultimately did not survive.

³¹ This is the relevant award for our above analysis of blockbuster awards.

D. Middleton v. Collins

Middleton v. Collins is a very unusual case for several reasons.³² While the case involved the largest (unadjusted for inflation) punitive damages award in history, the plaintiffs never expected to recover any part of the award. In fact, the plaintiffs filed suit primarily to protest inaction by the local prosecutor's office. In 1998, 13-year-old Don Collins attacked eight-year-old Robert Middleton. Collins doused Middleton with gasoline and set him on fire, and Middleton suffered severe burns over every part of his body except for one foot. Middleton received treatment at Shriners hospital, so he did not incur any medical expenses. However, he spent the rest of his life in and out of the hospital and required multiple surgeries, which physicians described as unimaginably painful for Middleton. Middleton died in 2010 from cancer caused by the multiple skin grafts he required. In a deathbed confession, Middleton said that in addition to the gasoline attack, Collins had raped him.

Local prosecutors declined to file charges against Collins for his attack on Middleton. However, Middleton's parents filed a civil suit in 2011 (after *State Farm* was decided) against Collins on behalf of themselves and their son's estate for battery and wrongful death among other claims. At trial, Collins did not retain counsel and filed a general denial pro se, but he never actually appeared in the courtroom—he was incarcerated for other crimes at the time. The court empaneled a jury, but the judge granted summary judgment on liability to the plaintiffs. The jury awarded each of Middleton's parents \$60 million and Middleton's estate \$250 million. The jury also awarded \$150 billion in punitive damages to the plaintiffs, which resulted in a punitive to compensatory damages ratio of approximately 405. However, the Middletons have publicly stated that they never expect to collect any part of this judgment and that they brought

³² This case was not reported in an official reporter, and the description here is based on Top 100 Verdicts of 2011 and other media.

the civil case primarily to stir a public outcry and pressure the state to prosecute Collins for his actions in criminal court. Therefore, while *Middleton* did involve a blockbuster punitive damages award, that award is deceptive because even though the award is real, it represents more of a statement than a true award of damages to the plaintiffs.

E. Discussion of the Four Largest Awards

Collectively, the litigation histories of the four largest awards provide useful insight into blockbuster awards. Even the succinct overview of the cases provided here illustrates the complexity of the litigation with three of the cases involving multiple appellate processes. More importantly, however, these four cases demonstrate that the initial blockbuster award is not always paid in its entirety. In all of the four largest cases except *Middleton*, the original blockbuster award was reduced or eliminated in subsequent litigation. Interestingly, *State Farm* played a role in the reductions in the two tobacco cases, so anecdotal evidence supports a meaningful *State Farm* effect. Even without *State Farm*, however, defendants often do not pay the full amount of the initial award. State law may, as it did in *Exxon Mobil Corp.*, require the reduction of an award, or the defendant may simply refuse to pay.³³

V. CONCLUSION

While punitive damages can serve a useful role in the judicial system in theory, evidence suggests that in some cases the awards may be very extreme. Blockbuster punitive damages awards exhibit an interesting type of unpredictability, as they follow a fat-tailed distribution. Predicting future awards when those awards follow a fat-tailed distribution can be particularly challenging if potential defendants assume that the awards follow more typical distributions, e.g.,

³³ Obviously, refusing to pay damages can create other problems for a defendant, but if the defendant is already incarcerated, as he was in *Middleton*, these problems are likely to prove less salient.

normal. Contrary to typical expectations about the size of awards, the largest award in a fat-tailed distribution can be several times larger than the next largest award. Therefore, predicting and planning for very large punitive damages awards can be very difficult for potential defendants.

The U.S. Supreme Court has consistently recognized the lack of predictability in punitive damages awards in general, not just for blockbuster awards, and has provided guidelines to foster restraint in such awards. The current constitutional doctrine on punitive damages, anchored by the *State Farm* decision, has been successful at reducing blockbuster punitive damages awards. *State Farm* altered the relationship between compensatory and punitive damages so that an additional dollar of compensatory damages generates less in additional punitive damages. This change in relationship is particularly important since the amount of compensatory damages awarded has consistently been shown to be the best indicator of the amount of punitive damages awarded in a given case.

Even more importantly, as we demonstrate in this paper, *State Farm* has effectively "thinned" the fat tail of the distribution of blockbuster awards. While award amounts decreased in general after *State Farm*, this effect was most prominent for the upper end of the distribution of awards. Thus, *State Farm* made blockbuster awards more predictable to the extent that a thinner tail allows defendants to better anticipate their liability. However, despite the noticeable shift in the inverse CDF of blockbuster awards after *State Farm*, these awards continue to follow a fat-tailed distribution. Future courts and policymakers alike will want to take note of this distribution and the implications it has for the predictability of awards as the laws surrounding punitive damages change in the future.

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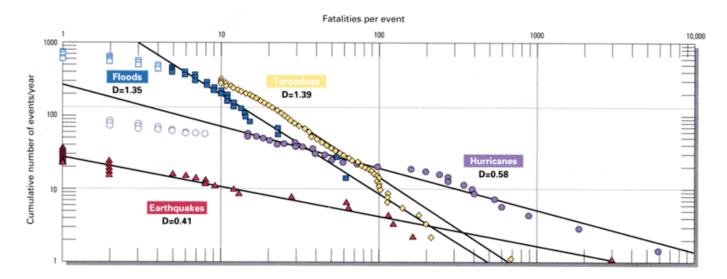
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Figures and Tables

Figure 1 Distributions of natural disasters



SOURCE: U.S. Geological Service Fact Sheet, "Natural Disasters – Forecasting Economic and Life Losses," http://pubs.usgs.gov/fs/natural-disasters/index.html.

Figure 2 Inverse cumulative distribution function of blockbuster punitive damages awards

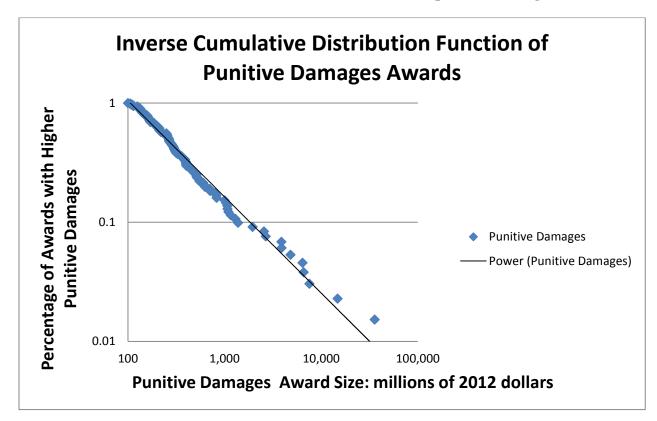


Figure 3 Histogram of blockbuster punitive damages awards (log amounts)

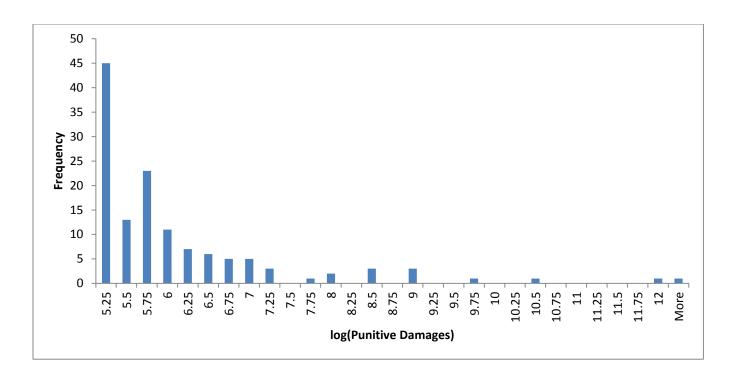


Figure 4 Blockbuster punitive damages (PD) awards over time, 1981 – 2012

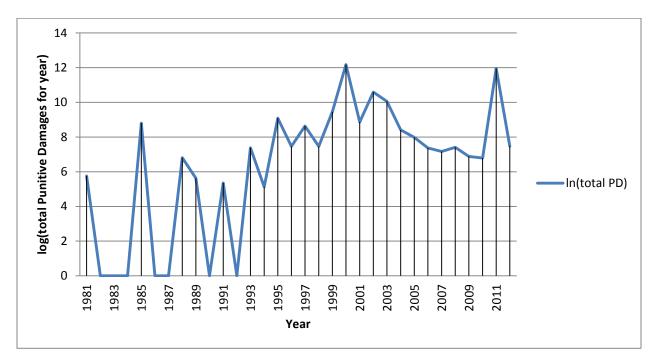
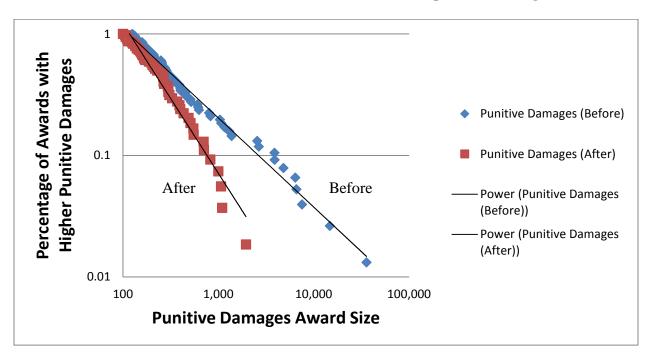


Figure 5 Inverse cumulative distribution functions of blockbuster punitive damages awards



Decile	Mean	Median	Sum	Minimum	Maximum	Number
First	118	114	1,770	100	133	15
Second	142	140	1,844	134	155	13
Third	164	165	2,132	158	171	13
Fourth	199	199	2,583	183	211	13
Fifth	245	252	3,188	212	267	13
Sixth	287	286	3,725	267	307	13
Seventh	367	370	4,765	312	402	13
Eighth	521	510	6,770	432	625	13
Ninth	997	1,036	12,958	700	1,378	13
Tenth	33,625	6,401	437,131	1,945	193,329	13

 Table 1

 Summary Statistics for Deciles of Blockbuster Punitive Damages Awards

NOTE: Damages are reported in millions of 2012 dollars and rounded to the nearest million.

	(1)	(2)
Variables	log(PD)	log(PD)
log(Compensatory Damages)	0.172***	0.204***
	(0.055)	(0.053)
log(Compensatory Damages)xState Farm	-0.036***	-0.032***
	(0.009)	(0.010)
Business and Individual Litigants		0.367*
		(0.198)
Automobile Industry		0.099
		(0.288)
Tobacco Industry		2.043**
		(0.869)
Finance, Investment, Insurance Industry		0.645***
		(0.207)
Energy, Chemical Industry		0.742***
		(0.250)
Pharmaceuticals, Health Industry		0.477**
		(0.209)
Violent Crime		0.690***
		(0.261)
Texas		-0.269
		(0.215)
California		0.216
		(0.235)
Bench Trial		0.097
		(0.292)
Constant	16.9***	15.1***
	(0.954)	(0.981)
R-squared	0.148	0.339

Table 2 OLS Regression Results of Log(Punitive Damages)

NOTE: Dependent variable is the natural log of punitive damages in all specifications. All awards are in 2012 dollars. Excluded industry category is other industry. All specifications exclude the *Garamendi* and *Chopourian* cases which involved no compensatory damages and the *Engle* and *Middleton* cases which involved the two largest PD Awards in the dataset. Robust standard errors are in parentheses. N = 128.

*** significant at 1% level; ** significant at 5% level; * significant at 10% level.

Variables	(1) 25th	(2) Median	(3) 75 <i>t</i> h
log(Compensatory Damages)	0.097***	0.180***	0.228***
	(0.022)	(0.040)	(0.080)
log(Compensatory Damages)xState Farm	-0.031***	-0.033***	-0.069***
	(0.009)	(0.013)	(0.021)
log(Compensatory Damages)xEngle	-0.013	-0.020	-0.044**
	(0.009)	(0.013)	(0.022)
Constant	17.6***	16.7***	16.9***
	(0.375)	(0.680)	(1.431)

Table 3Quantile Regression Results

NOTE: All columns report the quantile regression at the percentile noted in the column heading. Dependent variable is the natural log of punitive damages in all specifications. All specifications exclude the *Garamendi* and *Chopourian* cases which involved no compensatory damages. Robust standard errors are in parentheses. N = 130.

*** significant at 1% level; ** significant at 5% level; * significant at 10% level.

	(1)	(2)
Variables	I (<i>ratio</i> > 10)	I (<i>ratio</i> > 10)
log(Compensatory Damages)	-0.383***	-0.495***
	(0.067)	(0.010)
log(Compensatory Damages)xState Farm	-0.021***	-0.020**
	(0.007)	(0.009)
Business and Individual Litigants		0.138
		(0.173)
Automobile Industry		0.153
		(0.304)
Tobacco Industry		-0.079
		(0.170)
Finance, Investment, Insurance Industry		0.506***
		(0.159)
Energy, Chemical Industry		0.583***
		(0.152)
Pharmaceuticals, Health Industry		0.292
		(0.202)
Violent Crime		0.023
		(0.267)
Texas		-0.394***
		(0.100)
California		0.370*
		(0.210)
Bench Trial		0.018
		(0.264)

 Table 4

 Probit Regressions for Whether Awards Violated the Single Digit Ratio Cap

NOTE: Both columns report marginal effects of probit models. Dependent variable is an indicator for whether the PD to CD ratio exceeded 10. Excluded industry category is other industry. All specifications exclude the *Garamendi* and *Chopourian* cases which involved no compensatory damages, and the *Engle* and *Middleton* awards. Robust standard errors are in parentheses. N = 128.

*** significant at 1% level; ** significant at 5% level; * significant at 10% level.